

# Demographic Change, Household Structure and Income Inequality

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Sixth Winter School on Inequality and Collective Welfare Theory (IT6)  
Canazei, January 11, 2011

## Background

- OECD (2008): "*Growing unequal*": increasing inequality
- Question: what are the causes?
- OECD: high correlation between changing household composition and increasing inequality in West Germany (1985-2005): 88%
- this has led to a fierce policy debate in Germany
- however: result was a mistake! (Correct figure is 12%!)

# This talk

- How can we assess the question " *What drives rising inequality?*" ?
- Different methods:
  - ▶ Subgroup decomposition of inequality measures
  - ▶ Counterfactual reweighting techniques
    - ★ OECD: special case: shift-share analysis without control variables
- Examples for Germany:
  - ▶ A. Peichl, N. Pestel and H. Schneider (2010): *Does Size matter?: The Impact of Changes in Household Structure on Income Distribution in Germany*, CESifo Working Paper 3219
  - ▶ Biewen, M., Juhasz, A. (2010): *Understanding Rising Income Inequality in Germany*, IZA Discussion Paper No. 5062.

# Outline

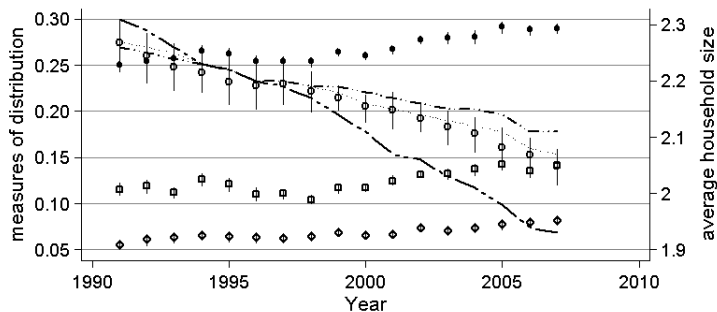
- 1 Why Germany?
- 2 Equivalence-weighting
- 3 Methodology
  - Subgroup Decomposition
  - Re-weighting procedure
- 4 Empirical Strategy
- 5 Results
- 6 BJ-Results
- 7 Conclusions
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## 1. Why Germany?

## Widening income gap, declining household size

- since reunification: inequality of disposable income distribution increased considerably (*Bach et al., 2009; Peichl et al., 2010*)
  - ▶ widening of market incomes / weakening bargaining power of unions?
  - ▶ structural change in household formation?
- observe sharp fall in average household size in Germany since early 1990s
  - ▶ second-lowest among OECD countries after Sweden
  - ▶ especially number of one- and two-person households increased
- link between trends: analysis of income distribution based on equivalent incomes
  - ▶ equivalence scales account for household structure (size and age)
  - ▶ i.e. changes in household structure *c.p.* influence income distribution

# Widening income gap, declining household size II



- Gini coefficient ( $I^{Gini}$ )
- ◊ Richness headcount ( $R^0$ )
- ◻ Poverty headcount ( $P^0$ )
- Household size (GSOEP)
- ⋯ Household size (Micro Census)
- only West (without Berlin)
- only East (with Berlin)

Source: German Micro Census and GSOEP, own calculations.  
Confidence intervals (95 per cent) based on 500 bootstrap replications.

## 2. Equivalence-weighting

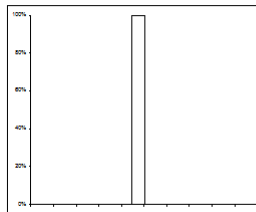


## Intuition

- economic well-being considered as *individual* experience
- **however:** individually received incomes *not* used for analysis of income distribution
- **reasons:**
  - ▶ dependent persons without resources for consumption
  - ▶ economies of scale in household consumption unconsidered
  - ▶ comparison of individuals irrespective of household size
- equivalent incomes serve as proxies for economic well-being

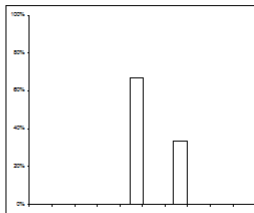
## Household Structure and Income Inequality

								
Income:	-	-	-	X	-	-	-	X
Equivalence scale:	.5	.3	.3	1	.5	.3	.3	1



Equivalence weighted income distribution

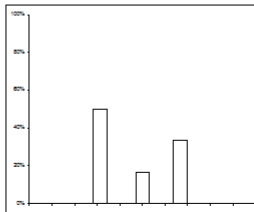
## Household Structure and Income Inequality



Equivalence weighted income distribution

## Household Structure and Income Inequality

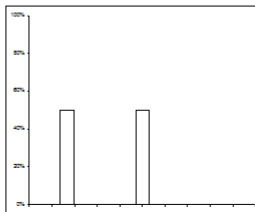
			
Income :	$X/2$ -   -	$X/2$	- $X$
Equivalence scale :	1   .3   .3	1	.5   1



Equivalence weighted income distribution

## Use of per capita incomes can increase the effect

			
Income :	$X/2$ -   -	$X/2$	-   X
Equivalence scale :	1   1   1	1	1   1



Equivalence weighted income distribution

### 3. Methodology

### 3.1 Subgroup Inequality Decomposition

- Shorrocks (1980, 1984); **Mookherjee/Shorrocks (1982)**
- **Jenkins (1995)**, Martin (2006)

## Mean logarithmic deviation

- most suitable: Generalized Entropy (GE) inequality measures
- decomposable for population subgroups  $k \in \{1, \dots, K\}$

$$l_0 = \frac{1}{n} \cdot \sum_{i=1}^n \ln \left( \frac{\bar{y}}{y_i} \right) \quad (1a)$$

$$= \sum_{k=1}^K v_k \cdot l_{0k} + \sum_{k=1}^K v_k \cdot \ln \left( \frac{\bar{y}}{\bar{y}_k} \right) \quad (1b)$$

- ▶  $y_i$ : equivalent individual income
- ▶  $\bar{y}$ : population mean income
- ▶  $v_k$ : proportion of population subgroup  $k$
- ▶  $l_{0k}/\bar{y}_k$ : inequality/mean income of subgroup  $k$



# Decomposition I

- decomposition of inequality *change* between periods  $t$  and  $t + 1$  (see *Mookherjee/Shorrocks, 1982*)

$$\Delta l_0 \approx \sum_{k=1}^K \bar{v}_k \cdot \Delta l_{0k} + \sum_{k=1}^K \bar{l}_{0k} \cdot \Delta v_k + \sum_{k=1}^K \left[ \bar{\lambda}_k - \overline{\ln(\lambda_k)} \right] \cdot \Delta v_k + \sum_{k=1}^K (\bar{\theta}_k - \bar{v}_k) \cdot \Delta \ln(\bar{y}_k) \quad (2)$$

- ▶  $\lambda_k = \bar{y}_k / \bar{y}$ : ratio of subgroup  $k$ 's mean income to total mean income
- ▶  $\theta_k = v_k \cdot \lambda_k$ : income ratio of group  $k$
- ▶ symbol with bar denotes average over periods  $t$  and  $t + 1$

## Decomposition II

$$\begin{aligned}
 \Delta I_0 \approx & \underbrace{\sum_{k=1}^K \bar{v}_k \cdot \Delta I_{0k}}_A + \underbrace{\sum_{k=1}^K \bar{I}_{0k} \cdot \Delta v_k}_B \\
 & + \underbrace{\sum_{k=1}^K [\bar{\lambda}_k - \overline{\ln(\lambda_k)}] \cdot \Delta v_k}_C + \underbrace{\sum_{k=1}^K (\bar{\theta}_k - \bar{v}_k) \cdot \Delta \ln(\bar{y}_k)}_D \quad (3)
 \end{aligned}$$

- *A*: change within population subgroups
- *B*: change in population composition on within inequality
- *C*: change in population composition on between inequality
- *D*: changes in population subgroup mean incomes
- **prior interest**: relative importance of *B* and *C* compared to  $\Delta I_0$

### 3.2 Re-weighting

- Di Nardo/Fortin/Lemieux, 1996; **Firpo/Fortin/Lemieux (2010)**
- **Hyslop/Maré, 2005**; Biewen/Juhasz (2010)

## Re-weighting procedure

- each individual described by vector  $(y, x, t)$ 
  - ▶ income  $y$ , household characteristics  $x$ , and time  $t$  vector
  - ▶ joint distribution  $F(y, x, t)$
- joint distribution of income and characteristics:  $F(y, x|t)$
- density of income at certain point in time:

$$f_t(y) = \int dF(y, x|t_{y,x} = t) = \int f(y|x, t_y = t)dF(x|t_x = t) \quad (4a)$$

$$\equiv f(y, t_y = t, t_x = t) \quad (4b)$$

- see *Di Nardo/Fortin/Lemieux, 1996; Hyslop/Maré, 2005*

## Re-weighting procedure II

- hypothetical counterfactual distribution:

$$f(y, t_y = 2007, t_x = 1991) = \int f(y|x, t_y = 2007) dF(x|t_x = 1991) \quad (5a)$$

$$= \int f(y|x, t_y = 2007) \psi_x(x) dF(x|t_x = 2007) \quad (5b)$$

- re-weighting function:

$$\psi_x(x) \equiv \frac{dF(x|t_x = 1991)}{dF(x|t_x = 2007)} \quad (6)$$

- counterfactual density can be estimated by weighted kernel methods

## 4. Empirical Strategy

## Data and income concept

- German Socio-Economic Panel Study (GSOEP)
  - ▶ panel survey of households and individuals in Germany conducted annually since 1984
  - ▶ weights allow representativeness for German population
- income concept:
  - ▶ pre and post fisc incomes
  - ▶ modified OECD equivalence scale
- 16 population groups:  
(No. of adults) X (No. of children) X (No. of earners)

$k$	adults	children	employed	$v_k^{1991}$	$\Delta v_k$	$\bar{y}_k^{0,post,1991}$	$\Delta \bar{y}_k^{0,post}$	$\hat{\beta}_{0,post}^{k,1991}$	$\Delta \hat{\beta}_{0,post}^k$	$\hat{\beta}_{0,pre}^{k,1991}$	$\Delta \hat{\beta}_{0,pre}^k$	$\hat{\rho}_{0,post}^{k,1991}$	$\Delta \hat{\rho}_{0,post}^k$	$\hat{R}_{0,post}^{k,1991}$	$\Delta \hat{R}_{0,post}^k$
1	1	no	0	0.090 (0.003)	0.011 (0.005)	14,102.35 (391.15)	1,718.73 (471.03)	0.125 (0.012)	0.029 (0.014)	1.216 (0.074)	-0.096 (0.086)	0.356 (0.020)	-0.032 (0.024)	0.019 (0.005)	0.018 (0.008)
2	1	no	1	0.067 (0.003)	0.031 (0.004)	21,660.89 (679.42)	48,648 (1,008).64	0.135 (0.019)	0.031 (0.030)	0.212 (0.026)	0.142 (0.037)	0.084 (0.011)	0.047 (0.015)	0.095 (0.015)	-0.012 (0.016)
3	1	yes	0	0.007 (0.001)	0.006 (0.002)	8,218.39 (566.93)	834.19 (635.94)	0.132 (0.025)	-0.077 (0.028)	0.437 (0.052)	0.635 (0.145)	0.732 (0.052)	-0.014 (0.062)	0.000 (0.000)	0.000 (0.000)
4	1	yes	1	0.021 (0.001)	0.004 (0.002)	13,726.20 (517.46)	-1,003.54 (543.96)	0.112 (0.011)	-0.032 (0.014)	0.218 (0.020)	0.191 (0.046)	0.323 (0.030)	0.046 (0.037)	0.035 (0.013)	-0.030 (0.013)
5	2	no	0	0.093 (0.003)	0.040 (0.005)	16,110.03 (370.28)	3,103.29 (509.75)	0.102 (0.011)	0.034 (0.014)	0.912 (0.047)	0.133 (0.062)	0.174 (0.012)	-0.030 (0.017)	0.034 (0.007)	0.030 (0.008)
6	2	no	1	0.072 (0.003)	0.014 (0.003)	20,820.02 (418.13)	3,177.36 (1,006).93	0.104 (0.008)	0.072 (0.025)	0.228 (0.020)	0.191 (0.037)	0.069 (0.008)	0.011 (0.014)	0.079 (0.012)	0.042 (0.016)
7	2	no	2	0.094 (0.003)	0.000 (0.004)	25,701.18 (418.21)	3,201.73 (527.45)	0.087 (0.007)	0.029 (0.009)	0.128 (0.009)	0.056 (0.014)	0.021 (0.004)	-0.001 (0.007)	0.157 (0.011)	0.065 (0.017)
8	2	yes	0	0.005 (0.001)	0.012 (0.001)	12,826.74 (601.46)	187.29 (857.11)	0.063 (0.013)	0.065 (0.020)	0.813 (0.189)	0.119 (0.215)	0.372 (0.056)	0.137 (0.066)	0.000 (0.000)	0.021 (0.008)
9	2	yes	1	0.137 (0.003)	-0.041 (0.004)	15,573.69 (146.15)	2,257.36 (245.92)	0.070 (0.003)	0.023 (0.006)	0.157 (0.009)	0.096 (0.017)	0.139 (0.009)	0.004 (0.014)	0.012 (0.003)	0.032 (0.007)
10	2	yes	$\geq 2$	0.185 (0.003)	-0.039 (0.005)	18,723.81 (157.71)	3,474.51 (346.61)	0.070 (0.003)	0.034 (0.006)	0.111 (0.005)	0.068 (0.010)	0.046 (0.003)	-0.001 (0.006)	0.045 (0.005)	0.045 (0.008)
11	$\geq 3$	no	0	0.006 (0.001)	0.002 (0.001)	18,819.59 (1,506.98)	-3,352.69 (1,718.37)	0.125 (0.015)	0.007 (0.023)	1.159 (0.148)	-0.403 (0.159)	0.279 (0.066)	0.064 (0.079)	0.103 (0.052)	-0.072 (0.053)
12	$\geq 3$	no	1	0.031 (0.002)	-0.003 (0.003)	19,508.20 (507.53)	359.316 (908.68)	0.079 (0.009)	0.055 (0.023)	0.264 (0.026)	0.088 (0.045)	0.090 (0.016)	0.044 (0.023)	0.031 (0.010)	0.019 (0.015)
13	$\geq 3$	no	$\geq 2$	0.113 (0.003)	-0.031 (0.004)	22,502.53 (217.405)	1,171.95 (388.02)	0.054 (0.002)	0.033 (0.005)	0.091 (0.003)	0.051 (0.009)	0.015 (0.002)	0.011 (0.005)	0.069 (0.006)	0.035 (0.011)
14	$\geq 3$	yes	0	0.000 (0.000)	0.003 (0.000)	11,030.41 (1,165).37	157.85 (1,386).60	0.020 (0.007)	0.018 (0.017)	0.839 (0.323)	-0.407 (0.322)	0.549 (0.262)	0.096 (0.275)	0.000 (0.000)	0.000 (0.000)
15	$\geq 3$	yes	1	0.015 (0.001)	0.004 (0.002)	16,383.19 (596.22)	-544.04 (758.52)	0.110 (0.012)	0.007 (0.013)	0.271 (0.028)	0.072 (0.039)	0.184 (0.027)	0.173 (0.046)	0.067 (0.016)	-0.052 (0.016)
16	$\geq 3$	yes	$\geq 2$	0.065 (0.002)	-0.012 (0.003)	18,302.44 (216.89)	811.29 (358.60)	0.066 (0.003)	0.003 (0.005)	0.102 (0.004)	0.031 (0.009)	0.063 (0.007)	-0.006 (0.014)	0.044 (0.007)	-0.011 (0.009)
Total	-	-	-	1.000 (0.000)	0.000 (0.000)	18,816.32 (106.66)	1,782.19 (162.88)	0.105 (0.002)	0.040 (0.004)	0.500 (0.010)	0.125 (0.016)	0.115 (0.003)	0.026 (0.005)	0.056 (0.002)	0.026 (0.003)



## 5. Results

# Inequality decomposition 1991–2007

income	region	$I_0^{1991}$	$I_0^{2007}$	$\Delta I_0$	A	B	C	D	$\frac{B+C}{\Delta I_0}$
household structure and employment status									
pre fisc	Germany	0.500 (0.010)	0.625 (0.011)	25.027 (3.542)	15.973 (2.274)	11.800 (1.211)	7.596 (0.973)	-10.148 (1.716)	77.500 (8.150)
	West	0.480 (0.012)	0.558 (0.012)	16.284 (4.042)	15.892 (2.658)	7.982 (1.210)	5.542 (1.048)	-12.870 (1.836)	83.052 (16.407)
	East	0.514 (0.022)	0.872 (0.024)	69.567 (8.524)	15.711 (3.743)	28.931 (3.154)	23.860 (3.097)	-0.584 (3.691)	75.885 (5.311)
post fisc	Germany	0.105 (0.002)	0.144 (0.004)	37.755 (4.463)	28.917 (3.991)	5.354 (0.682)	3.024 (0.586)	0.560 (1.415)	22.189 (2.851)
	West	0.104 (0.003)	0.149 (0.004)	42.990 (5.268)	35.679 (4.635)	4.689 (0.694)	2.145 (0.656)	0.564 (1.508)	15.896 (2.248)
	East	0.070 (0.002)	0.097 (0.003)	38.801 (6.022)	44.055 (4.886)	-0.731 (1.639)	7.239 (1.938)	-16.178 (2.479)	16.773 (8.656)

# Inequality decomposition 1991–2007

$\frac{B+C}{\Delta I_0}$  for different equivalence scales  $ES = (\theta_1 + \theta_2 \cdot N_A + \theta_3 \cdot N_C)^\gamma$  :

	$\theta_1 = \theta_2 = 0.5$						$\theta_1 = 0; \theta_2 = 1$					
	$\theta_3 = 0.3$		$\theta_3 = 0.5$		$\theta_3 = 1$		$\theta_3 = 0.3$		$\theta_3 = 0.5$		$\theta_3 = 1$	
income	$\gamma = 0.5$	$\gamma = 1$	$\gamma = 0.5$	$\gamma = 1$	$\gamma = 0.5$	$\gamma = 1$	$\gamma = 0.5$	$\gamma = 1$	$\gamma = 0.5$	$\gamma = 1$	$\gamma = 0.5$	$\gamma = 1$
	household structure and employment status											
pre fisc	79.143 (6.336)	77.500 (5.798)	79.319 (6.391)	78.139 (5.972)	78.931 (6.315)	76.762 (5.740)	78.497 (6.064)	77.941 (5.618)	78.698 (6.084)	78.591 (5.747)	78.307 (6.080)	77.322 (5.594)
post fisc	23.259 (2.285)	22.189 (2.482)	23.353 (2.570)	22.853 (3.212)	22.797 (2.027)	20.054 (1.925)	21.658 (2.145)	24.296 (2.888)	22.264 (2.373)	26.476 (3.498)	20.751 (1.958)	21.075 (2.471)

# Re-weighting results 1991-2007

measure	pre fisc			post fisc		
	$\Delta^{act}$	$\Delta^{rew}$	$\frac{\Delta^{act} - \Delta^{rew}}{\Delta^{act}}$	$\Delta^{act}$	$\Delta^{rew}$	$\frac{\Delta^{act} - \Delta^{rew}}{\Delta^{act}}$
$I_{Gini}$	18.39 (1.44)	9.16 (1.26)	50.21 (3.21)	16.14 (1.65)	12.45 (1.53)	22.85 (2.54)
$I_0$	25.03 (3.59)	4.97 (2.92)	80.14 (9.42)	37.76 (4.46)	28.82 (3.91)	23.67 (2.54)
$I_1$	39.97 (5.45)	20.69 (4.24)	48.24 (3.90)	54.24 (10.34)	43.11 (8.47)	20.51 (2.75)
$I_2$	107.12 (37.28)	66.74 (26.45)	37.70 (4.11)	187.16 (81.27)	148.65 (65.29)	20.58 (3.14)
post fisc incomes						
	poverty			richness		
$P_0/R_0$	22.60 (5.11)	10.65 (4.52)	52.87 (13.06)	46.62 (7.20)	40.26 (7.24)	13.64 (4.58)
$P_1/R_3$	36.36 (7.74)	21.08 (6.95)	42.03 (9.28)	65.75 (9.69)	56.79 (9.54)	13.63 (2.93)
$P_2/R_1$	47.24 (11.48)	29.44 (10.22)	37.68 (10.65)	76.06 (11.54)	65.90 (11.36)	13.36 (2.85)

## Summary and discussion

- proportion of “demographic effect” much larger for pre fisc incomes
- tax-benefit system seems to compensate for changing household structure at bottom of distribution
- however, no causal relationship: tax-benefit system itself might have enforced demographic trends
- results of subgroup decomposition in line with those of a counterfactual re-weighting analysis (without further controls!)

## 6. Biewen / Juhasz

# Approach

- Re-weighting à la Hyslop/Maré controlling for various characteristics
  - ▶ advantage: several distributional statistics can be computed
  - ▶ advantage: can control for other characteristics
  - ▶ disadvantage: path-dependence
- GSOEP 1999+2000 vs. 2005+2006 (pooled data!)
- only look at post fisc (disposable) income

# Results

- increase in inequality can be explained by
  - ▶ changes in employment outcomes and market returns
  - ▶ and changes in the tax system.
- Changes in household structures and other household characteristics seem to have played a much smaller role.
- However: several issues with the analysis! (data, method, weights, policy modelling ...)

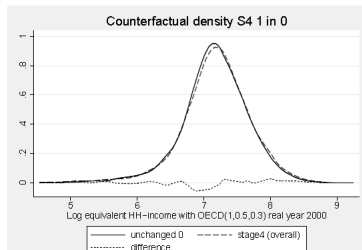
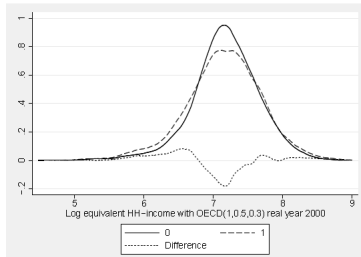


# Results II

Overall change in density from 1999/2000 ('period 0') to 2005/2006 ('period 1')

Counterfactual income distribution if only market returns are changed (dashed line) vs.

factual distribution (bold line).



## Results III

Table 4 – Exact decomposition of inequality increase

	Results of sequential decomposition attributable to										
	Household Structure		Socio-economic attributes		Employment outcomes		Return on attributes		Tax system		Residual
	(1)	(2)	(3)	(4)	(5)						
<b>p5010</b>	7.39	(2.82)	5.96	(3.26)	30.48	(8.76)	23.62	(16.49)	24.30	(5.64)	8.25
<b>p7525</b>	6.80	(2.29)	3.42	(2.79)	22.54	(5.33)	14.15	(10.66)	19.08	(3.22)	33.30
<b>p9010</b>	8.93	(2.45)	6.04	(2.96)	30.16	(7.11)	29.61	(13.03)	20.59	(3.77)	4.67
<b>p9050</b>	13.47	(4.92)	6.77	(7.24)	30.80	(12.70)	41.89	(24.98)	10.62	(8.59)	-3.55
<b>Cv</b>	8.20	(2.21)	4.66	(3.04)	16.96	(4.19)	22.76	(7.80)	20.92	(5.24)	26.50
<b>Theil</b>	8.33	(2.24)	5.07	(2.70)	19.92	(4.54)	31.41	(9.80)	19.88	(4.66)	15.36
<b>Mld</b>	3.90	(2.23)	5.81	(2.70)	23.30	(5.43)	28.85	(12.1)	19.64	(4.65)	18.47
<b>Gini</b>	5.31	(2.44)	5.54	(2.79)	23.17	(4.99)	17.71	(10.91)	17.77	(4.71)	30.48
<b>Fgt0</b>	7.72	(2.58)	5.34	(2.73)	26.67	(6.64)	20.23	(12.34)	19.81	(3.97)	20.24
<b>Fgt1</b>	4.03	(2.54)	8.21	(3.79)	30.40	(9.07)	39.38	(17.19)	23.09	(5.01)	-5.11

Source: GSOEP, own calculations. The numbers in parentheses are bootstrap standard errors which correctly take into account the longitudinal sample design and the clustering of individuals in households.

## 7. Conclusion

# Conclusions

- changing household composition associated with widening income gap
- **but:** share of 15% for post fisc incomes only (*for inequality*)
  - ▶ much lower than reported by OECD
  - ▶ other more important driving forces
  - ▶ human capital? bargaining power of unions? → future research
- statements on income distribution must be differentiated
  - ▶ important to analyze different reasons for a growing income gap
  - ▶ complex interactions between income distribution, demographic trends (household formation), and tax-benefit system
  
- Detailed policy decomposition: see Bargain et al. (2011)

# Thank you for your attention!

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## FGT measures

- well-known poverty measure  $P_\alpha$  (Foster et al., 1984)
- richness measure  $R_\beta$  (Peichl et al., 2008)
- decomposable for population subgroups

$$P_\alpha(y; z) = \frac{1}{n} \cdot \sum_{i=1}^n \left( \frac{z - y_i}{z} \right)^\alpha \cdot \mathbf{1}_{y_i < z} \quad (7)$$

$$R_\beta(y; \rho) = \frac{1}{n} \cdot \sum_{i=1}^n \left[ 1 - \left( \frac{\rho}{y_i} \right)^\beta \right] \cdot \mathbf{1}_{y_i > \rho} \quad (8)$$

- ▶  $z$ : poverty line,  $\rho$ : richness line
- ▶  $\alpha$ : parameter for poverty aversion,  $\beta$ : parameter for sensitiveness to (intense) richness

# Decomposition

$$\Delta P_\alpha = \underbrace{\sum_{k=1}^K \bar{v}_k \cdot \Delta P_{\alpha,k}}_A + \underbrace{\sum_{k=1}^K \bar{P}_{\alpha,k} \cdot \Delta v_k}_B \quad (9)$$

$$\Delta R_\beta = \underbrace{\sum_{k=1}^K \bar{v}_k \cdot \Delta R_{\beta,k}}_A + \underbrace{\sum_{k=1}^K \bar{R}_{\beta,k} \cdot \Delta v_k}_B \quad (10)$$

- $A$ : change in level of group poverty/richness
- $B$ : changes in composition of population
- **prior interest**: relative importance of  $B$  relative to  $\Delta P_\alpha$  and  $\Delta R_\beta$