

Unambiguous Comparison of Intersecting Distribution Functions

Rolf Aaberge, *Statistics Norway*

Tarjei Havnes, *Univ of Oslo, ESOP & Statistics Norway*

Magne Mogstad, *Univ College London, Statistics Norway & ESOP*

January 9, 2012

How do we compare intersecting distribution functions?

Important issue in both policy work, descriptive analysis and causal inference:

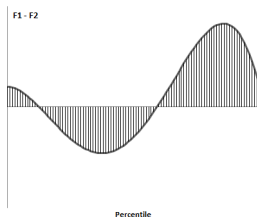
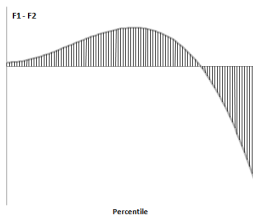
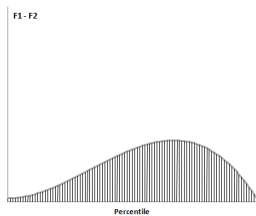
- 1 statistical offices and gov agencies compare distribution functions across countries, subgroups and time
- 2 descriptive research compares distributions of earnings, income, consumption and wealth to evaluate economic welfare
- 3 growing interest in econometrics in how to estimate the counterfactual outcome distribution
 - yet little attention has been devoted to how to compare counterfactual and actual outcome distributions

Suppose we want to rank the actual and counterfactual distributions, F_1 and F_2

- Straightforward with 1st or 2nd-degree dominance
 - but many empirical applications require weaker criteria

Theoretical literature: Offers higher order dominance criteria

Empirical literature: Tends to use parametric social welfare function

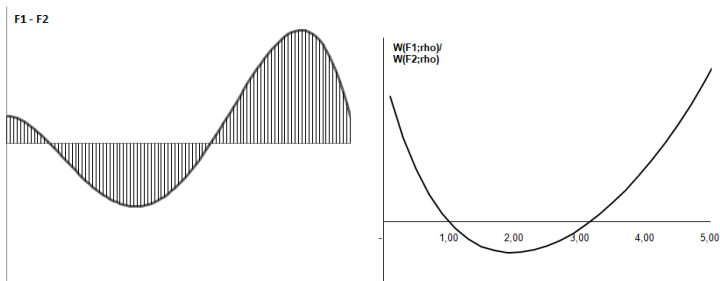


General dominance criteria: Hard to interpret and justify

- Rely on assumptions about third and higher order derivatives (see e.g Atkinson, 2003)

Parametric social welfare functions:

- Conclusion rests on more or less arbitrary parameter choice (and functional form)
- Ranking is non-monotonic in inequality aversion
 - An example: $W(F) = \int \frac{y^{1-\rho}}{1-\rho} dF(y)$, $\rho \in [0, \infty)$
 $\rho = 0$: ineq. neutral, $\rho = 1$: log, $\rho \rightarrow \infty$: mini-max



Aim: Proposes a general framework to *unambiguously* compare *any set* of distributions functions in an *economically interpretable* way

- ① Social welfare functions and 2nd-degree dominance
- ② Social welfare functions and 3rd-degree upward and downward dominance
- ③ Social welfare functions and i th-degree upward and downward dominance
- ④ Parametric subfamilies
 - Upward: Gini family
 - Downward: Lorenz family
- ⑤ Asymptotic theory
- ⑥ Application

The general family of social welfare function

We will rely on the general family of rank-dependent measures of social welfare introduced by Yaari (1987,1988)

$$W_P(F) = \int_0^1 P'(t)F^{-1}(t)dt,$$

The weighting function P' is the derivative of a preference function that is a member of the following the set of preference functions:

$$\mathcal{P} = \{P : P'(t) > 0 \text{ and } P''(t) < 0 \\ \text{for all } t \in (0, 1), P(0) = P'(1) = 0, P(1) = 1\}$$

The general family of social welfare function

We will rely on the general family of rank-dependent measures of social welfare introduced by Yaari (1987,1988)

$$W_P(F) = \int_0^1 P'(t)F^{-1}(t)dt,$$

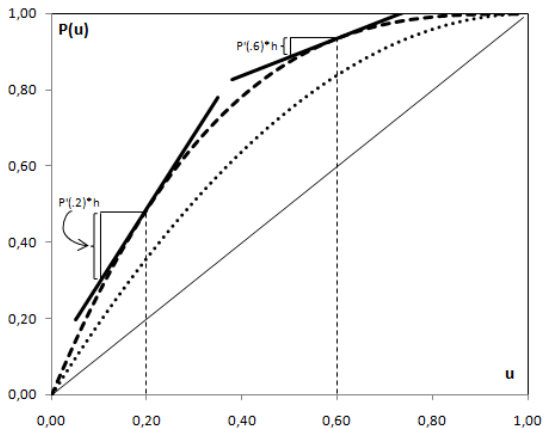
The weighting function P' is the derivative of a preference function that is a member of the following the set of preference functions:

$$\mathcal{P} = \{P : P'(t) > 0 \text{ and } P''(t) < 0 \\ \text{for all } t \in (0, 1), P(0) = P'(1) = 0, P(1) = 1\}$$

- W_P preserves 1st-degree dom, since $P'(t) > 0$, and
- W_P preserves 2nd-degree dom (and Pigou-Dalton), since $P''(t) < 0$
- $W_P \leq \mu_F$, and $W_P = \mu_F$ iff F is the egalitarian distribution

The preference function: Examples

$P(t)$ reveals the inequality aversion profile of the social planner



Normative justification of the general family

The normative justification of W_P can be made in terms of a

(a) Theory for ranking distribution functions:

- With basic ordering and continuity assumptions, the dual independence axiom characterizes W_P (Yaari, 1988)

(b) Value judgement of the trade-off between the mean and (in)equality in the distributions (Ebert, 1987; Aaberge, 2001)

$$W_P = \mu_F [1 - J_P(F)]$$

where μ_F is the mean of F and

$J_P(F)$ is the family of rank-dependent measures of inequality aggregating the P' -weighted Lorenz curve of F

If we choose

$$P_{1k}(t) = 1 - (1 - t)^{k-1}, \quad k > 2$$

then W_P is equal to the extended Gini family of social welfare functions (Donaldson and Weymark, 1980)

$$W_{G_k} = \mu [1 - G_k(F)], \quad k > 2$$

where

- $G_k(F)$ is the extended Gini family of inequality measures
- $G_3(F)$ is the Gini coefficient and $W_{G_2} = \mu$
- Note that $\{\mu, W_{G_i}(F) : i = 3, 4, \dots\}$ uniquely determines the distribution function F (Aaberge, 2000)

If we instead choose

$$P_{2k}(t) = \frac{(k-1)t - t^{k-1}}{k-2}, \quad k > 2$$

then W_P is the Lorenz family of social welfare functions (Aaberge, 2000)

$$W_{D_k} = \mu [1 - D_k(F)], \quad k > 2$$

where

- $D_k(F)$ is the Lorenz family of inequality measures
- $D_3(F)$ is the Gini coefficient
- Note that $\{\mu, W_{D_i}(F) : i = 3, 4, \dots\}$ uniquely determines the distribution function F (Aaberge, 2000)

Third degree upward dominance

Note that second degree inverse stochastic dominance is defined by

$$\Lambda_F^2(u) \equiv \int_0^u F^{-1}(t) dt, \quad u \in [0, 1]$$

To define third degree upward inverse stochastic dominance, we use the notation

$$\Lambda_F^3(u) \equiv \int_0^u \Lambda_F^2(t) dt = \int_0^u (u-t) F^{-1}(t) dt, \quad u \in [0, 1]$$

Definition

A distribution F_1 is said to *third degree upward inverse stochastic dominate* a distribution F_0 if and only if

$$\Lambda_{F_1}^3(u) \geq \Lambda_{F_0}^3(u) \text{ for all } u \in [0, 1]$$

and the inequality holds strictly for some $u \in (0, 1)$.

Proposition

Let F_1 and F_0 be members of F . Then the following statements are equivalent:

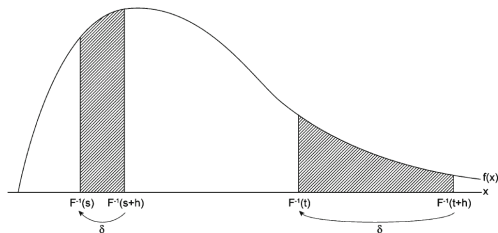
(i) F_1 third degree upward inverse stochastic dominates F_0

(ii) $\mu_{F_1}(u) (1 - G_3(u; F_1)) \geq \mu_{F_0}(u) (1 - G_3(u; F_0))$

for all $u \in [0, 1]$ and the inequality holds strictly for some $u \in (0, 1)$.

where:

- $\mu_F(u)$ is the quantile-specific lower tail mean
- $G_3(u; F)$ is the quantile-specific lower tail Gini coefficient
- $\mu_F(u) (1 - G_3(u; F))$ is the quantile-specific lower tail Gini social welfare function



$\Delta_s W_P(\delta, h)$: change in W_P of a fixed progressive transfer δ from an individual with rank $s + h$ to an individual with rank s .

$$\Delta_{st}^1 W_P(\delta, h) \equiv \Delta_s W_P(\delta, h) - \Delta_t W_P(\delta, h).$$

Definition

(Zoli, 1999; Aaberge, 2000, 2009) W_P satisfies the principle of first degree downside positional transfer sensitivity (DPTS) if and only if

$$\Delta_{st}^1 W_P(\delta, h) > 0, \quad \text{when } s < t.$$

Let \mathcal{P}_3 be the family of preference functions defined by

$$\mathcal{P}_3 = \left\{ P \in \mathcal{P} : P'''(t) > 0, \right\}$$

Theorem

Let F_1 and F_0 be members of F . Then the following statements are equivalent.

- (i) F_1 third-degree upward inverse stochastic dominates F_0
- (ii) $W_P(F_1) > W_P(F_0)$ for all $P \in \mathcal{P}_3$
- (iii) $W_P(F_1) > W_P(F_0)$ for all $P \in \mathcal{P}$ where W_P satisfies first-degree DPTS

⇒ (i) and (ii): least-restrictive set of social welfare functions that unambiguously rank in accordance with 3-UID

⇒ (i) and (iii): normative justification for 3-UID

Upward vs. downward dominance

Upward dominance criteria justified through DPTS

- More sensitive to changes in the lower part of the distribution

Issues with upward dominance criteria:

- Prone to measurement error in the lower tail
- Changes in the upper part may be viewed as more important
 - Long vs. short transfers
 - Upper part is the focus: Test scores, top income, etc.

Upward vs. downward dominance

Upward dominance criteria justified through DPTS

- More sensitive to changes in the lower part of the distribution

Issues with upward dominance criteria:

- Prone to measurement error in the lower tail
- Changes in the upper part may be viewed as more important
 - Long vs. short transfers
 - Upper part is the focus: Test scores, top income, etc.

We propose a complementary sequence of dominance criteria:

⇒ Downward inverse stochastic dominance

- More sensitive to changes in the upper part of the distribution

Sequences coincide at 2nd-degree dom. ⇒ both obey Pigou-Dalton

Third-degree downward dominance

The criteria of 3rd-order downward inverse dominance aggregates $\Lambda_F^2(u)$ from above (rather than from below). Let

$$\tilde{\Lambda}_F^3(u) \equiv \int_u^1 \Lambda_F^2(t) dt = (1-u)\mu_F - \int_u^1 (t-u)F^{-1}(t)dt, \quad u \in [0, 1]$$

Definition

A distribution F_1 is said to *third degree downward inverse stochastic dominate* a distribution F_0 if and only if

$$\tilde{\Lambda}_{F_1}^3(u) \geq \tilde{\Lambda}_{F_0}^3(u) \text{ for all } u \in [0, 1]$$

and the inequality holds strictly for some $u \in (0, 1)$.

Proposition

Let F_1 and F_2 be members of F . Then the following statements are equivalent:

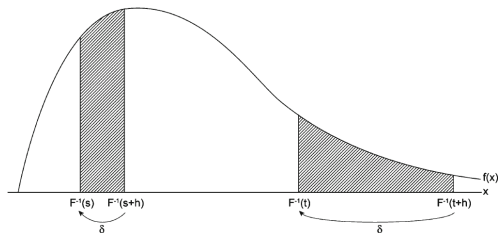
(i) F_1 third degree downward inverse stochastic dominates F_2

(ii) $\tilde{\mu}_{F_1}(u)(1 - D_3(u; F_1)) \geq \tilde{\mu}_{F_2}(u)(1 - D_3(u; F_2))$

for all $u \in [0, 1]$ and the inequality holds strictly for some $u \in (0, 1)$

where:

- $\tilde{\mu}_F(u)$ is the quantile-specific upper tail mean
- $D_3(u; F)$ is the quantile-specific upper tail Gini coefficient
- $\tilde{\mu}_F(u)(1 - D_3(u; F))$ is the quantile-specific upper tail Gini social welfare function



$\Delta_s W_P(\delta, h)$: change in W_P of a fixed progressive transfer δ from an individual with rank $s + h$ to an individual with rank s .

$$\Delta_{st}^1 W_P(\delta, h) \equiv \Delta_s W_P(\delta, h) - \Delta_t W_P(\delta, h).$$

Definition

(Aaberge, 2009) W_P satisfies the principle of first degree upside positional transfer sensitivity (UPTS) if and only if

$$\Delta_{st}^1 W_P(\delta, h) < 0, \quad \text{when } s < t.$$

Let $\tilde{\mathcal{P}}_3$ be the family of preference functions defined by

$$\tilde{\mathcal{P}}_3 = \left\{ P \in \mathcal{P} : P'''(t) < 0 \right\}.$$

Theorem

Let F_1 and F_0 be members of F . Then the following statements are equivalent.

- (i) F_1 third-degree downward inverse stochastic dominates F_0
- (ii) $W_P(F_1) > W_P(F_0)$ for all $P \in \tilde{\mathcal{P}}_3$
- (iii) $W_P(F_1) > W_P(F_0)$ for all $P \in \mathcal{P}$ where W_P satisfies first-degree UPTS

\Rightarrow (i) and (ii): least-restrictive set of social welfare functions that unambiguously rank in accordance with 3-DID

\Rightarrow (i) and (iii): normative justification for 3-DID

To define upward inverse stochastic dominance of degree i , we use the notation

$$\begin{aligned}\Lambda_F^i(u) &= \int_0^u \Lambda_F^{i-1}(t) dt = \frac{1}{(i-3)!} \int_0^u (u-t)^{i-3} \Lambda_F^2(t) dt \\ &= \frac{1}{(i-2)!} \int_0^u (u-t)^{i-2} F^{-1}(t) dt, \quad i > 2\end{aligned}$$

Definition

A distribution F_1 is said to i th degree upward inverse stochastic dominate F_0 for $i > 2$ if and only if

$$\Lambda_{F_1}^i(u) \geq \Lambda_{F_0}^i(u) \text{ for all } u \in [0, 1]$$

and the inequality holds strictly for some $u \in (0, 1)$.

Proposition

Let F_0 and F_1 be members of \mathcal{F} . Then for $i = 3, 4, \dots$ the following statements are equivalent:

(i) F_1 i^{th} degree upward inverse stochastic dominates F_0

(ii) $\mu_{F_1}(u) (1 - G_i(u; F_1)) \geq \mu_{F_0}(u) (1 - G_i(u; F_0))$.

for all $u \in [0, 1]$ and the inequality holds strictly for some $u \in (0, 1)$.

where:

- $G_i(u; F)$ is the quantile-specific lower tail i^{th} member of the Gini family of inequality measures
- $\mu_F(u) (1 - G_i(u; F))$ is the quantile-specific lower tail i^{th} member of the Gini family of social welfare functions

The family of preference functions \mathcal{P}_i is defined by

$$\mathcal{P}_i = \left\{ P \in \mathcal{P} : (-1)^{i-1} P^{(i)}(t) > 0 \text{ with } P^{(i)} \text{ continuous on } (0, 1) \right. \\ \left. \text{and } (-1)^{i-1} P^{(j)}(1) \geq 0 \text{ for all } j = 3, 4, \dots, i-1 \right\}$$

where $P^{(i)}$ denote the i th degree derivative of P .

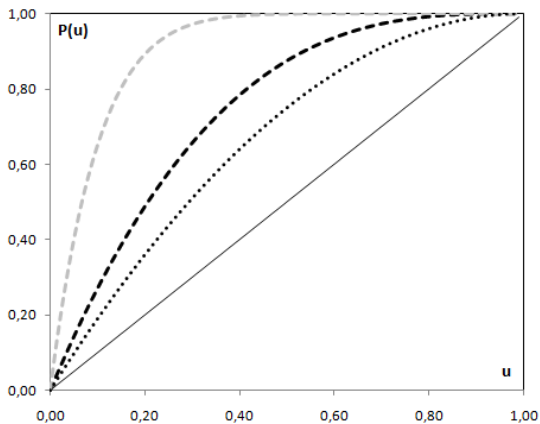
Theorem

Let F_1 and F_0 be members of \mathcal{F} . Then for $i = 3, 4, \dots$ the following statements are equivalent,

- (i) F_1 i th degree upward inverse stochastic dominates F_0
- (ii) $W_P(F_1) > W_P(F_0)$ for all $P \in \mathcal{P}_i$
- (iii) $W_P(F_1) > W_P(F_0)$ for all $P \in \mathcal{P}$ where W_P satisfies DPTS of degree $i - 2$

- \Rightarrow (i) and (ii): least-restrictive set of social welfare functions that unambiguously rank in accordance with i th-degree UID
- \Rightarrow (i) and (iii): normative justification for i th-degree UID

Upward dominance: The weighting function



Downward dominance of i^{th} degree

To define downward inverse stochastic dominance of degree i , we use the notation

$$\begin{aligned}\tilde{\Lambda}_F^i(u) &= \int_u^1 \tilde{\Lambda}_F^{i-1}(u) = \frac{1}{(i-3)!} \int_u^1 (t-u)^{i-3} \Lambda_F^2(t) dt = \\ &\frac{1}{(i-2)!} \left[(1-u)^{i-2} \mu_F - \int_u^1 (t-u)^{i-2} F^{-1}(t) dt \right] \\ &i = 3, 4, \dots\end{aligned}$$

Definition

A distribution F_1 is said to i th degree downward inverse stochastic dominate F_0 for $i > 2$ if and only if

$$\tilde{\Lambda}_{F_1}^i(u) \geq \tilde{\Lambda}_{F_0}^i(u) \text{ for all } u \in [0, 1]$$

and the inequality holds strictly for some $u \in (0, 1)$.

Proposition

Let F_0 and F_1 be members of \mathcal{F} . Then for $i = 3, 4, \dots$ the following statements are equivalent:

(i) F_1 i^{th} degree downward inverse stochastic dominates F_0

(ii) $\tilde{\mu}_{F_1}(u)(1 - D_i(u; F_1)) \geq \tilde{\mu}_{F_0}(u)(1 - D_i(u; F_0))$.

for all $u \in [0, 1]$ and the inequality holds strictly for some $u \in (0, 1)$.

where:

- $D_i(u; F)$ is the quantile-specific upper tail i^{th} member of the Lorenz family of inequality measures
- $\tilde{\mu}_F(u)(1 - D_i(u; F))$ is the quantile-specific upper tail i^{th} member of the Lorenz family of social welfare functions

The family of preference functions $\tilde{\mathcal{P}}_i$ is defined by

$$\tilde{\mathcal{P}}_i = \left\{ P \in \mathcal{P} : P^{(i)}(t) < 0 \text{ with } P^{(j)} \text{ continuous on } (0,1) \quad (1) \right. \\ \left. \text{and } P^{(j)}(0) \leq 0 \text{ for all } j = 3, 4, \dots, i-1 \right\}$$

where $P^{(i)}$ denote the i th degree derivative of P .

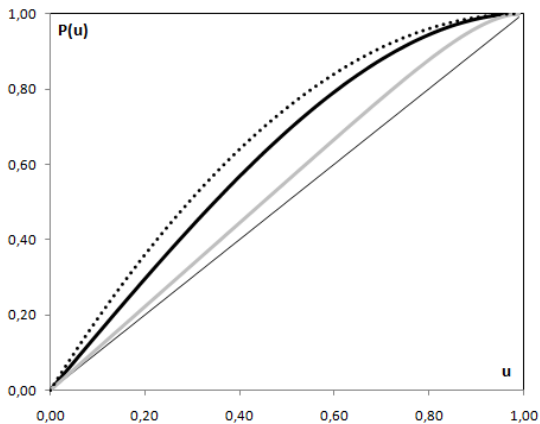
Theorem

Let F_1 and F_0 be members of \mathcal{F} . Then for $i = 3, 4 \dots$ the following statements are equivalent

- (i) F_1 i th degree downward inverse stochastic dominates F_0
- (ii) $W_P(F_1) > W_P(F_0)$ for all $P \in \tilde{\mathcal{P}}_i$
- (iii) $W_P(F_1) > W_P(F_0)$ for all $P \in \mathcal{P}$ where W_P satisfies UPTS of degree $i - 2$

- \Rightarrow (i) and (ii): least-restrictive set of social welfare functions that unambiguously rank in accordance with i th-degree DID
- \Rightarrow (i) and (iii): normative justification for i th-degree DID

Downward dominance: The weighting function



Limit of the sequences of dominance

As i goes to infinity, we get from the definitions of upward and downward dominance:

$$(i-1)! \Lambda^i(u) \rightarrow \begin{cases} 0, & 0 \leq u < 1 \\ F^{-1}(0+), & u = 1 \end{cases}$$
$$(i-2)! \tilde{\Lambda}^i(u) \rightarrow \begin{cases} \int \mu - F^{-1}(1-) dt, & u = 0 \\ 0, & 0 < u \leq 1 \end{cases}$$

where $F^{-1}(0+)$ and $F^{-1}(1-)$ denote the lowest and highest income in F

- Limit upward dominance: Social welfare function corresponding to the (Rawlsian) maximin criterion
- Limit downward dominance: Social welfare function approaches the utilitarian criterion

Proposition

Let F_1 and F_0 be members of F . Then for $i = 3, 4, \dots$

(i) F_1 i^{th} degree upward inverse stochastic dominates F_0
implies

(ii) $W_{G_k}(F_1) > W_{G_k}(F_0)$ for $k > i$

Remark. The extended Gini family of social welfare functions has the following properties.

(i) W_{G_i} preserves upward inverse stochastic dominance of degree $< i$

(ii) W_{G_i} obeys the Pigou-Dalton principle of transfers

(iii) W_{G_i} obeys the principles of DPTS up to and including

$(i - 2)$ th-degree for $i = 3, 4, \dots$

(iv) The sequence $\{W_{G_i}\}$ approaches μ_F as $i \rightarrow 2$

(v) The sequence $\{W_{G_i}\}$ approaches the Rawlsian maximin criterion as $i \rightarrow \infty$.

Proposition

Let F_1 and F_0 be members of F . Then $i = 3, 4..$

(i) F_1 i th degree downward inverse stochastic dominates F_0
implies

(ii) $W_{D_k}(F_1) > W_{D_k}(F_0)$ for $k > i$

Remark. The extended Lorenz family of social welfare functions has the following properties,

(i) W_{D_i} preserves downward inverse stochastic dominance of degree $< i$

(ii) W_{D_i} obeys the Pigou-Dalton principle of transfers.

(iii) W_{D_i} obeys the principles of UPTS up to and including $(i - 2)$ th-degree.

(iv) The sequence $\{W_{D_i}\}$ approaches μ_F as $i \rightarrow \infty$

(v) The sequence $\{i(W_{D_i} - \mu_F)\}$ approaches $\mu_F - \int_0^1 F^{-1}(t) dt$ as $i \rightarrow \infty$

Weights at quantiles relative to median

Quantile:	$0+$	$.05$	$.30$	$.70$	$.95$	$1-$
Panel (a): Gini social welfare function (upward)						
$i \rightarrow 2$	1.00	1.00	1.00	1.00	1.00	1.00
$i = 3$	2.00	1.90	1.40	0.60	0.10	$0+$
$i = 4$	4.00	3.61	1.96	0.36	0.01	$0+$
$i = 5$	8.00	6.86	2.74	0.22	0.00	$0+$
$i = 6$	16.00	13.03	3.84	0.13	0.00	$0+$
$i \rightarrow \infty$	∞	0	0	0	0	0
Panel (b): Lorenz social welfare function (downward)						
$i = 3$	2.00	1.90	1.40	0.60	0.10	$0+$
$i = 4$	1.33	1.33	1.21	0.68	0.13	$0+$
$i = 5$	1.14	1.14	1.11	0.75	0.16	$0+$
$i = 6$	1.07	1.07	1.06	0.81	0.20	$0+$
$i \rightarrow \infty$	1	1	1	1	1	$0+$

1) Since F_n is a consistent estimator of F

- $\Lambda_{F_n}^i(u)$ and $\tilde{\Lambda}_{F_n}^i(u)$ are consistent estimators of $\Lambda_F^i(u)$ and $\tilde{\Lambda}_F^i(u)$

2) The asymptotic properties of $\Lambda_{F_n}^i(u)$ and $\tilde{\Lambda}_{F_n}^i(u)$ can be obtained by

- considering the limiting distribution of the empirical processes

$$Y_n^i(u) = \sqrt{n} [\Lambda_{F_n}^i(u) - \Lambda_F^i(u)]$$

$$\tilde{Y}_n^i(u) = \sqrt{n} [\tilde{\Lambda}_{F_n}^i(u) - \tilde{\Lambda}_F^i(u)]$$

We can then show that $\tilde{Y}_n^i(u)$ and $Y_n^i(u)$

- converge to a Gaussian process and thus are asymptotically normally distributed

Theorem

Let $W_0(t)$ denote a Brownian bridge on $[0, 1]$. Suppose that F has a continuous nonzero derivative f on $[a, b]$. Then $Y_n^i(u)$ converges in distribution to the processes

$$Y^i(u) = \frac{1}{(i-2)!} \int_0^u (u-t)^{i-2} \frac{W_0(t)}{f(F^{-1}(t))} dt$$

which has the same probability distribution as the Gaussian process $\sum_{j=1}^{\infty} h_j(u) Z_j$, where $h_j(u)$ is given by

$$h_j(u) = \frac{1}{(i-2)!} \left[\frac{\sqrt{2}}{j\pi} \int_0^u (u-t)^{i-2} \frac{\sin(j\pi t)}{f(F^{-1}(t))} dt \right]$$

and Z_1, Z_2, \dots are independent $N(0, 1)$ -variables.

Theorem

Let $W_0(t)$ denote a Brownian bridge on $[0, 1]$. Suppose that F has a continuous nonzero derivative f on $[a, b]$. Then $\tilde{Y}_n^i(u)$ converges in distribution to the processes

$$\tilde{Y}^i(u) = \frac{1}{(i-2)!} \left[(1-u)^{i-2} \int_0^1 \frac{W_0(t)}{f(F^{-1}(t))} dt - \int_u^1 (t-u)^{i-2} \frac{W_0(t)}{f(F^{-1}(t))} dt \right]$$

which has the same probability distribution as the Gaussian process $\sum_{j=1}^{\infty} \tilde{h}_j(u) Z_j$, where $\tilde{h}_j(u)$ is given by

$$\tilde{h}_j(u) = \frac{1}{(i-2)!} \frac{\sqrt{2}}{j\pi} \left[(1-u)^{i-2} \int_0^1 \frac{\sin(j\pi t)}{f(F^{-1}(t))} dt - \int_u^1 (t-u)^{i-2} \frac{\sin(j\pi t)}{f(F^{-1}(t))} dt \right]$$

and Z_1, Z_2, \dots are independent $N(0, 1)$ -variables.

Application: Jobs First field experiment

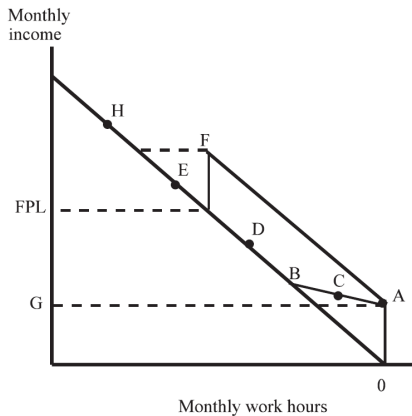
We apply our framework to the Jobs First program Apr 96–Dec 00, analyzed in Bitler et al. (2005, AER)

- Random assignment to Jobs First or AFDC
- Two counties in Connecticut: New Haven and Manchester
- Sample of about 4803 welfare recipients

Key features of Job First program:

- Expanded earnings disregard
- Introduced 21 month time limit

Application: Jobs First – Budget constraint



$AB = AFDC$ $AF = \text{Jobs First}$

Application: Jobs First – Estimation

We use QTE-estimates from Bitler et al. (2008).

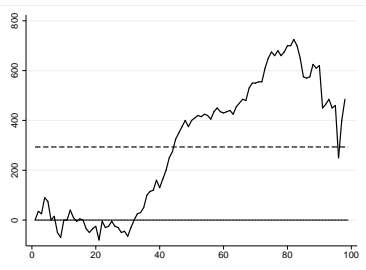
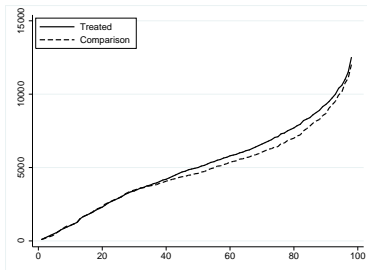
- Compares the quantiles of the treatment and control distribution: $\Delta_q = F_1^{-1}(q) - F_0^{-1}(q)$

Outcomes: Total income

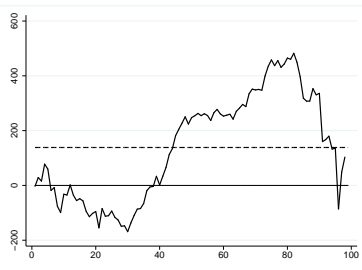
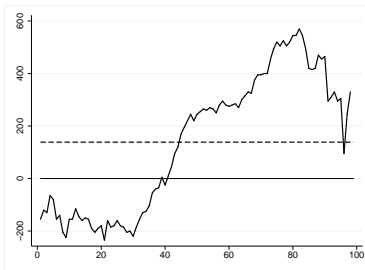
Financial costs:

- Job First: Higher cash transfers, admin costs, and operating costs
 - Assess gains and losses with\without balanced budget

QTE: Averaged income q1-q16



is



Dom. and soc welfare: Averaged income q1–q16

	Upward dominance			Downward dominance		
	<i>No tax</i>	<i>Lump sum</i>	<i>Prop. Tax</i>	<i>No tax</i>	<i>Lump sum</i>	<i>Prop. Tax</i>
Inverse stochastic dominance						
Dom.	4	4	231	3	3	3
Distr.	1	0	0	1	1	1
ΔW_p	8.8%	-6.2%	N/A	10.9%	0.6%	3.9%
$W(F_0)$	\$341	\$341	N/A	\$742	\$742	\$742
Social welfare weights relative to median						
p(.05)	3.61	3.61	7E+63	1.90	1.90	1.90
p(.30)	1.96	1.96	3E+33	1.40	1.40	1.40
p(.70)	0.36	0.36	2E-51	0.60	0.60	0.60
p(.95)	0.01	0.01	1E-229	0.10	0.10	0.10

We characterize the relationship between dominance criteria and two nested subfamilies of least restrictive social welfare functions

- higher-order UID = stronger downside inequality aversion
- higher-order DID = stronger upside inequality aversion
 - Useful to unambiguously say whether F_1 is better than F_0

We characterize the relationship between dominance criteria and two nested subfamilies of least restrictive social welfare functions

- higher-order UID = stronger downside inequality aversion
- higher-order DID = stronger upside inequality aversion
 - Useful to unambiguously say whether F_1 is better than F_0

We then derive parametric subfamilies of these social welfare functions that are easily implementable

- UID \Rightarrow Gini family, W_{G_i}
- DID \Rightarrow Lorenz family, W_{D_i}
 - Can inform on how much better F_1 is than F_0
 - Clarifies the dominance criterion in terms of observable social welfare weights

We characterize the relationship between dominance criteria and two nested subfamilies of least restrictive social welfare functions

- higher-order UID = stronger downside inequality aversion
- higher-order DID = stronger upside inequality aversion
 - Useful to unambiguously say whether F_1 is better than F_0

We then derive parametric subfamilies of these social welfare functions that are easily implementable

- UID \Rightarrow Gini family, W_{G_i}
- DID \Rightarrow Lorenz family, W_{D_i}
 - Can inform on how much better F_1 is than F_0
 - Clarifies the dominance criterion in terms of observable social welfare weights

We illustrate the usefulness of the framework by applying to an experimental policy intervention